

AMS 205B - ASSIGNMENT 3 DRAFT

Question 1 (CB 6.10)

By definition, a sufficient statistic is not complete if \exists a function $g(\cdot)$ such that $E(g(T(X))) = 0 \quad \forall \theta$ and $g(T(X)) \neq 0$.

The minimal sufficient statistic as found in class is $(X_{(1)}, X_{(n)})$.

Looking at the distribution of $X_{(1)}, X_{(n)}$ we have:

$$\begin{aligned} P(X_{(n)} \leq x) &= 1 - (1 - P(X \leq x))^n \\ &= 1 - (1 - (x - \theta))^n \end{aligned}$$

$$\begin{aligned} P(X_{(1)} \leq x) &= P(X \leq x)^n \\ &= (x - \theta)^n \end{aligned}$$

$$z = 1 - x + \theta \quad x = 1 + \theta - z$$

Taking the expected values, we have:

$$\bullet E(X_{(n)}) = \int_{\theta}^{\theta+1} n(1-x+\theta)^{n-1} x \, dx = \int_0^1 n z^{n-1} (1+\theta-z) \, dz = z^n (1+\theta) - \frac{n}{n+1} z^{n+1} \Big|_0^1$$

$$\bullet E(X_{(1)}) = \int_{\theta}^{\theta+1} n(x-\theta)^{n-1} x \, dx = \int_0^1 n z^{n-1} (z+\theta) \, dz = \frac{n z^{n+1}}{n+1} + z^n \theta \Big|_0^1 = \frac{n}{n+1} + \theta$$

$$\therefore g(T(x)) = X_{(n)} - X_{(1)} + \frac{2n}{n+1} - 1 = X_{(n)} - X_{(1)} + \frac{n-1}{n+1} \neq 0$$

$$\text{but } E(g(T(X))) = 0 \quad \forall \theta.$$

Hence, $T(X) = (X_{(1)}, X_{(n)})$ is not complete statistic for θ .

Question 5

Part a.

$$E(T(X_1, \dots, X_{n+1})) = 1 \cdot P_p(\sum_{i=1}^n X_i > X_{n+1}) + 0 = P_p(\sum_{i=1}^n X_i > X_{n+1}) = h(p)$$

$\therefore T(X_1, \dots, X_{n+1})$ is an unbiased estimator of $h(p)$. ~~not 25/100~~

Part b.

First we find a sufficient statistic for p . We have shown in class that $\sum_{i=1}^{n+1} X_i$ is a complete sufficient statistic.

Now using Rao-Blackwell, we can find that UMVUE = $E(T | \sum_{i=1}^n X_i = s)$

$$E(T | \sum_{i=1}^{n+1} X_i = s) = \frac{P(T=1 | \sum_{i=1}^n X_i = s)}{P(\sum_{i=1}^{n+1} X_i = s)}$$

- If $s=0$, then $P(\sum_{i=1}^n X_i > X_{n+1}, \sum_{i=1}^{n+1} X_i = 0) = 0$
- If $s \geq 3$, then $P(\sum_{i=1}^n X_i > X_{n+1}, \sum_{i=1}^{n+1} X_i \geq 3) = P(\sum_{i=1}^{n+1} X_i \geq 3)$
since $\sum_{i=1}^n X_i \geq 2$ and $X_{n+1} = \{0, 1\}$

- If $s=1$, then

$$P(\sum_{i=1}^n X_i > X_{n+1}, \sum_{i=1}^{n+1} X_i = 1) = n p (1-p)^{n-1} (1-p)$$

$\begin{matrix} \text{one of } X_i = 1 \\ i=1, \dots, n \\ X_{n+1} = 0 \end{matrix}$

$$P(\sum_{i=1}^{n+1} X_i = 1) = (n+1) p (1-p)^n$$

- If $s=2$, then

$$P(\sum_{i=1}^n X_i > X_{n+1}, \sum_{i=1}^{n+1} X_i = 2) = \binom{n}{2} p^2 (1-p)^{n-2} (1-p)$$

$\begin{matrix} \text{two } X_i = 1 \\ i=1, \dots, n \\ X_{n+1} = 0 \end{matrix}$

$$P(\sum_{i=1}^{n+1} X_i = 2) = \frac{n(n+1)}{2} p^2 (1-p)^{n-1}$$

In sum, the UMVUE is: $t = \sum_{i=1}^{n+1} X_i$

$$\phi(t) = \begin{cases} 0 & \text{if } t=0 \\ n(n+1)^{-1} & \text{if } t=1 \\ (n-1)(n+1)^{-1} & \text{if } t=2 \\ 1 & \text{if } t \geq 3 \end{cases}$$

Question 2

Answer